Efficient nonparametric inference for discretely observed compound Poisson processes

Alberto Coca Cabrero* CCA, University of Cambridge

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9th June 2016

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- Continuous observations
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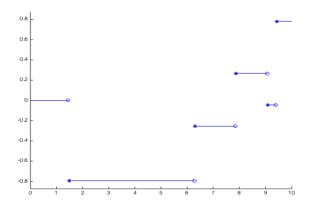
Construction and properties

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- R Efficient nonparametric inference for discretely observed CPPs

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Applications and literature

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Setting and estimators

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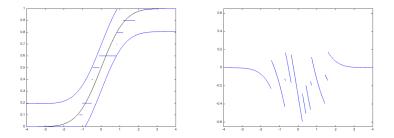
How to make inference on λ and F? Conditioned on $N_T = n$,

- λ can be estimated (parametrically) using the maximum likelihood estimator (asymptotic normality), and
- F can be estimated (nonparametrically) by the *empirical* distribution function F_n(x) := ¹/_n ∑ⁿ_{k=1} 1_{(-∞,x]}(X_k), x ∈ ℝ.

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Properties of $F_n(x) := \frac{1}{n} \sum_{k=1}^n \mathbb{1}_{(-\infty,x]}(X_k)$

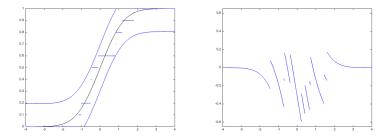


Left figure: Black: F = N(0,1); Blue: $F_n, n = 5, 10, \dots, 250$ and $F \pm \max |\widehat{F}_n - F|$

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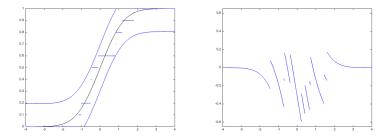
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Donsker's theorem:

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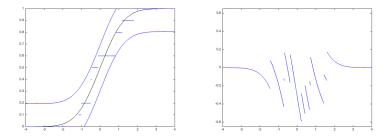
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Donsker's theorem: $\sqrt{n}(F_n - F) \rightarrow^{\mathcal{D}} \mathbb{G}_F$ in $L^{\infty}(\mathbb{R})$ as $n \rightarrow \infty$, where \mathbb{G}_F is the mean-zero Gaussian process on \mathbb{R} with covariance function $E[\mathbb{G}_F(x)\mathbb{G}_F(y)] = F(x \wedge y) - F(x)F(y)$.

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Discrete observations

Setting

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Setting

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• Instead, we observe $C_{\Delta}, \ldots, C_{n\Delta}$ for some $\Delta > 0$ and $n \in \mathbb{N}$.

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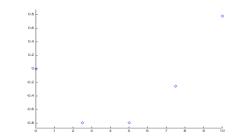


Figure: $C_{\Delta}, \ldots, C_{n\Delta}$, $\Delta = 2.5$ and n = 4 ($\lambda = 0.5, F = N(0, 1)$)

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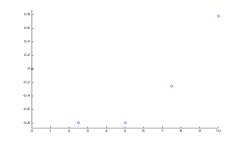


Figure: $C_{\Delta}, ..., C_{n\Delta}, \Delta = 2.5$ and n = 4 ($\lambda = 0.5, F = N(0, 1)$)

How can we infer λ and F with such incomplete information?

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• we are effectively observing a random variable X corrupted by a sum of a random number of independent copies of itself.

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Question: Is (nonparametric) $1/\sqrt{n}$ -consistent and asymptotically efficient estimation of *F* even possible?

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Question: Is (nonparametric) $1/\sqrt{n}$ -consistent and asymptotically efficient estimation of F even possible? **Answer:** Yes! We resort to the spectral approach to find a heuristic reason and to construct such an estimator.

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Then, the characteristic function of each increment (\sim the noise) is

$$\varphi(u) := \mathcal{F}[P](u) := E[e^{iuY}] = e^{\Delta(\mathcal{F}[d\mathcal{N}](u) - \lambda)}, \quad u \in \mathbb{R}.$$

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where Log is the distinguished logarithm

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$$\frac{1}{\Delta}f_{x}(y)\mathcal{F}^{-1}[\operatorname{Log}\varphi](dy) = f_{x}(y)(\mathcal{N}(dy) - \lambda\delta_{0}(dy)),$$

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where $f_{x,n} := \mathbb{1}_{(-\infty,x]} \mathbb{1}_{[-H_n,H_n] \setminus (-\varepsilon_n,\varepsilon_n)}$, $\varepsilon_n, H_n^{-1} \to 0$ as $n \to \infty$

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$$\varphi(u) := \mathcal{F}[P](u) := E[e^{iuY}] = e^{\Delta(\mathcal{F}[d\mathcal{N}](u) - \lambda)}, \quad u \in \mathbb{R}.$$

Due to $\|\mathcal{F}[d\mathcal{N}]\|_{L^{\infty}} \leq \lambda$, $\inf_{u \in \mathbb{R}} |\varphi(u)| \geq e^{-2\Delta\lambda} > 0$ so no *ill-posedness*! Furthermore, an estimator for \mathcal{N} can be constructed from it:

$$\mathcal{N}_n(x) := \frac{1}{\Delta} \int_{\mathbb{R}} f_{x,n}(y) \mathcal{F}^{-1}[\log \varphi_n](dy),$$

where $f_{x,n} := \mathbb{1}_{(-\infty,x]} \mathbb{1}_{[-H_n,H_n]\setminus(-\varepsilon_n,\varepsilon_n)}$, $\varepsilon_n, H_n^{-1} \to 0$ as $n \to \infty$, $\varphi_n(u) := \frac{1}{n} \sum_{k=1}^n e^{iuY_k}$ is the empirical characteristic function of the increments

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where $f_{x,n} := \mathbb{1}_{(-\infty,x]} \mathbb{1}_{[-H_n,H_n]\setminus(-\varepsilon_n,\varepsilon_n)}$, $\varepsilon_n, H_n^{-1} \to 0$ as $n \to \infty$, $\varphi_n(u) := \frac{1}{n} \sum_{k=1}^n e^{iuY_k}$ is the empirical characteristic function of the increments and $K_{h_n} := \frac{1}{h_n} K\left(\frac{\cdot}{h_n}\right)$, with K a band-limited kernel function and $h_n \to 0$.

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Due to $\lim_{x\to\infty} \mathcal{N}(x) = \lambda$, $\lambda_n := \mathcal{N}_n(\infty)$ and $\widehat{F}_n := \mathcal{N}_n/\lambda_n$.

Continuous observations Discrete observations

The spectral approach: $\widehat{F}_n := \mathcal{N}_n / \lambda_n$

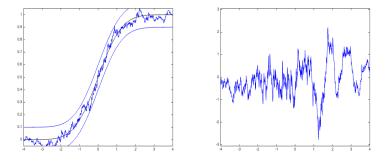
How does \widehat{F}_n and its \sqrt{n} -fluctuations about F look?

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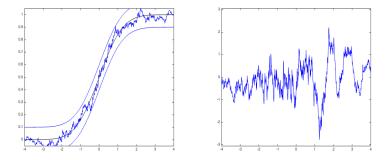


Left figure: Black: F = N(0,1); Blue: $\hat{F}_n, n = 10, 20, \dots, 500$ and $F \pm \max|\hat{F}_n - F|$

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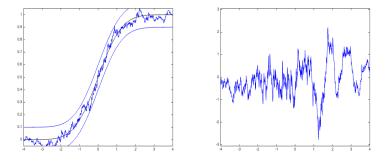
Left figure: Black: F = N(0,1); Blue: $\hat{F}_n, n = 10, 20, ..., 500$ and $F \pm \max|\hat{F}_n - F|$ Right figure: $\sqrt{n}(\hat{F}_n - F), n = 10, 20, ..., 500 (\lambda = 3, \Delta = 1)$

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The spectral approach: $F_n := \mathcal{N}_n / \lambda_n$

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Left figure: Black: F = N(0,1); Blue: $\hat{F}_n, n = 10, 20, ..., 500$ and $F \pm \max|\hat{F}_n - F|$ Right figure: $\sqrt{n}(\hat{F}_n - F), n = 10, 20, ..., 500$ ($\lambda = 3, \Delta = 1$) Can we show an analogue of Donsker's theorem for \hat{F}_n ?

We make no assumptions on $\lambda > 0$ and on F we assume

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- $|F(x)-F(y)| \lesssim |x-y|^{lpha}$ for all $x, y \in \mathbb{R}$ and some $lpha \in (0,1]$, and
- $\int_{\mathbb{R}} \log^{\beta} \left(\max\{|x|, e\} \right) F(dx) < \infty \text{ for some } \beta > 2.$

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- $|K(x)| \lesssim (1+|x|)^{-\eta}$ for some $\eta > 2$.

Continuous observations Discrete observations

A Donsker theorem for discrete observations

Theorem (Coca (2015))

Take $h_n \sim \exp(-n^{\vartheta_h})$, $\varepsilon_n \sim \exp(-n^{\vartheta_{\varepsilon}})$ and $H_n \sim \exp(n^{\vartheta_H})$ with $0 < \vartheta_{\varepsilon}, \vartheta_H < \vartheta_h < 1/4$.

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$$\sqrt{n}(\lambda_n - \lambda) \rightarrow^d N(0, \sigma^2),$$

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where $\sigma^2 := G(\mathbb{1}_{\mathbb{R} \setminus \{0\}}, \mathbb{1}_{\mathbb{R} \setminus \{0\}})$, and

$$\sqrt{n}\left(\widehat{F}_n-F\right)\to^{\mathcal{D}}\widehat{\mathbb{G}}_{\mathcal{N}}$$
 in $L^{\infty}(\mathbb{R}),$

where $\widehat{\mathbb{G}}_{\mathcal{N}}$ is a zero-mean Gaussian process on \mathbb{R} with covariance function $\sum_{x,y} := G(f_x, f_y)$, with $f_x := \lambda^{-1} (\mathbb{1}_{(-\infty,x]} - F(x)) \mathbb{1}_{\mathbb{R} \setminus \{0\}}$.

A Donsker theorem for discrete observations: Covariance

The covariances are defined through

$$G(g_1,g_2):=\frac{1}{\Delta^2}\int_{\mathbb{R}}(g_1*\mathcal{F}^{-1}[1/\varphi(-\cdot)](x))(g_2*\mathcal{F}^{-1}[1/\varphi(-\cdot)](x))P(dx),$$

where * denotes the convolution operation, $\mathcal{F}^{-1}\left[\varphi^{-1}(-\cdot)\right]$ is a finite signed measure on $\mathbb R$ satisfying

$$\mathcal{F}^{-1}\left[arphi^{-1}(-\cdot)
ight] = e^{\Delta\lambda} \sum_{k=0}^{\infty} ar{
u}^{*k} rac{(-\Delta)^k}{k!},$$

with $\bar{\nu}(A) := \nu(-A)$ for all $A \subseteq \mathbb{R}$ Borel measurable, $\bar{\nu}^{*k}$ is the *k*-fold convolution of $\bar{\nu}$ and $\bar{\nu}^{*0} = \delta_0$, and the probability measure *P* has the well-known representation (see Remark 27.3 in Sato (1999))

$$P = e^{-\Delta\lambda} \sum_{k=0}^{\infty} \nu^{*k} \frac{\Delta^k}{k!}.$$

Efficient nonparametric inference for discretely observed CPPs

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A Donsker theorem for discrete observations: Conclusions

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- (Optimal) confidence bands and goodness-of-fit tests for F can be constructed boostrap techniques such as substituting f, F, φ and P by their empirical counterparts in G;
- ΔλΣ_{x,y} = F(x ∧ y) F(x)F(y) + O(Δλ) so when Δλ is small classical Donsker's theorem is recovered and these procedures can be approximated by analogues independent of F, φ and P.

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A Donsker theorem for discrete observations: Proof

• Write $\sqrt{n}(\widehat{F}_n - F) = \lambda_n^{-1}\sqrt{n}((\mathcal{N}_n - \mathcal{N}) + F(\lambda - \lambda_n))$. Therefore we need joint convergence of λ_n and \mathcal{N}_n and, in particular, convergence of the latter:

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- Decompose $\sqrt{n}(N_n N)$ into a stochastic and a bias term and show uniform negligibility of the latter;

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- Write √n (F̂_n − F) = λ_n⁻¹ √n ((N_n − N) + F (λ − λ_n)). Therefore we need joint convergence of λ_n and N_n and, in particular, convergence of the latter:
- Decompose $\sqrt{n} (N_n N)$ into a stochastic and a bias term and show uniform negligibility of the latter;
- Split the stochastic term into its linear (in (P_n − P), P_n:= ¹/_n∑ⁿ_{k=1}δ_{Y_k}) and nonlinear part and show uniform negligibility of the latter by controlling sup_{|u|≤h_n⁻¹} |φ_n(u) − φ(u)|;

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- The linear term is $(P_n P) \psi_{x,n}$, where $Q \psi = \int_{\mathbb{R}} \psi \, dQ$ and $\psi_{x,n} := f_{x,n} * \mathcal{F}^{-1}[1/\varphi(-\cdot)] * K_{h_n}$. This is an empirical process indexed by a class of functions changing with n so, to show it is P-Donsker, we use the following theorem.

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Theorem (Theorem 2.11.23 in van der Vaart and Wellner (1996))

For each n, let $\Psi_n := \{\psi_{x,n} : x \in \mathbb{R}\}$ be a class of measurable functions indexed by a totally bounded semimetric space (\mathbb{R}, ρ) .

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$$\int_0^{\delta_n} \sqrt{\log N_{[]}(\epsilon \|\Psi_n\|_{L^2(P)}, \Psi_n, L^2(P))} \, d\epsilon, \sup_{\rho(x,y) < \delta_n} P(\psi_{x,n} - \psi_{y,n})^2 \to 0$$

where $\|\psi\|_{L^2(P)} = (\int_{\mathbb{R}} |\psi|^2 P)^{1/2}$.

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$$\int_0^{\sigma_n} \sqrt{\log N_{[]}(\epsilon \|\Psi_n\|_{L^2(P)}, \Psi_n, L^2(P))} \, d\epsilon, \sup_{\rho(x,y) < \delta_n} P(\psi_{x,n} - \psi_{y,n})^2 \to 0$$

where $\|\psi\|_{L^2(P)} = (\int_{\mathbb{R}} |\psi|^2 P)^{1/2}$. Then $(\sqrt{n}(P_n - P)\psi_{x,n})_{x \in \mathbb{R}}$ is asymptotically tight in $L^{\infty}(\mathbb{R})$ and converges in distribution to a tight Gaussian process provided the sequence of covariance functions $P\psi_{x,n}\psi_{y,n} - P\psi_{x,n}P\psi_{y,n}$ converges pointwise on $\mathbb{R} \times \mathbb{R}$.

In Coca (2015) we also show

 a Donsker theorem for N. The limit process, of Brownian motion-type, is different from that of F, of Brownian bridge-type. The former provides optimal inference procedures for the CPP as a whole and gives insight into efficiency issues;

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Furthermore, the results can be extended to X being multidimensional and to noisy (unknown but observed noise) and nonequispaced discrete observations. Future manuscript?

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Outline



2 Estimation of compound Poisson processes

- Continuous observations
- Discrete observations



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Thanks for your attention!

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