

Asymptotic distribution of the diameter of a random elliptical cloud

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Menu of the day

Random cloud

- \mathbb{X} is a random vector in \mathbb{R}^d with $d \geq 1$ fixed
- $n \geq 1$ is the size of the cloud
- $\{\mathbb{X}_i\}_{1 \leq i \leq n}$ are independent vectors distributed as \mathbb{X}

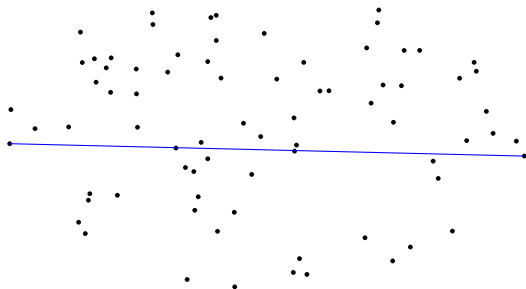


Menu of the day

Diameter of the cloud

- $\|\cdot\|$ is the Euclidean norm in \mathbb{R}^d

$$D_n := \max_{1 \leq i < j \leq n} \|\mathbb{X}_i - \mathbb{X}_j\|$$



Menu of the day

What is the asymptotic distribution of D_n when $n \rightarrow \infty$?

Answer ?

- For special cases
- Depends on the distribution of \mathbb{X}

Dichotomy

- Distributions supported by a bounded set
 - Distributions 'approximately' uniform
 - Geometry of the support
- Distributions supported by an unbounded set
 - Spherically symmetric distributions

History : bounded support

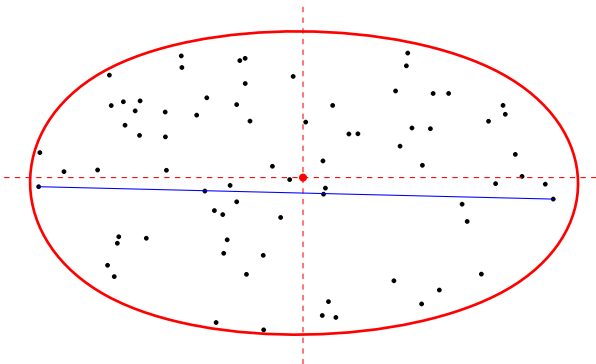
- Uniform distribution supported by special planar sets (excluding balls or ellipsoids) : *Appel, Najim and Russo* (2002)
- Distributions with support included in the unit d -ball (including uniform in the d -ball, in the d -sphere, in spherical sectors) : *Mayer and Molchanov* (2007)
- Distributions supported by a polytope (included uniform or non-uniform in square, uniform in regular polygons, uniform in the unit d -cube) : *Lao* (2010)
- Distributions supported by a d -ellipsoid : *Schrempp* (2016)

History : unbounded support

- Spherically symmetric normal distribution : *Matthews and Rukhin* (1993)
- Spherically symmetric Kotz distribution : *Henze and Klein* (1996)
- Power-tailed spherically decomposable distributions : *Henze and Lao* (2010)
- Spherically symmetric distributions : *Jammalamadaka and Janson* (2015)
 - Open question : elliptically symmetric distributions?

A naive question ...

Where are the points which can achieve the diameter?



... a naive answer

$$M_n := \max_{1 \leq i \leq n} \|\mathbb{X}_i\|$$

*D_n is achieved for a pair of diametrically opposed points
each of them realizing M_n*

If you believe in this, you need :

- To localize the vectors with large norms
- To control the asymptotic distribution of M_n

Precisely :

- Distribution of $\|\mathbb{X}\|$?
- Distribution of $\frac{1}{\|\mathbb{X}\|} \mathbb{X}$ conditional on $\|\mathbb{X}\|$ is large ?

Today's ingredients

Elliptical distribution : $\mathbb{X} = R\Lambda\mathbb{U}$

where

- $\mathbb{U} = (U_1, \dots, U_d)$ is uniform on the unit sphere \mathcal{S}^{d-1}
- Λ is an invertible $d \times d$ matrix
- R is a positive random variable independent of \mathbb{U}

In addition :

R is in the max domain of attraction of the Gumbel distribution

Today's ingredients

R is in the max domain of attraction of the Gumbel distribution

There exists a differentiable function $\psi_R : (0, \infty) \rightarrow (0, \infty)$ such that

$$\lim_{x \rightarrow \infty} \frac{\mathbb{P}(R > x + t\psi_R(x))}{\mathbb{P}(R > x)} = e^{-t}$$

locally uniformly with respect to $t \in \mathbb{R}$

Such a function ψ_R satisfies :

$$\lim_{x \rightarrow \infty} \frac{\psi_R(x + t\psi_R(x))}{\psi_R(x)} = 1 ; \quad \lim_{x \rightarrow \infty} \psi_R'(x) = 0 ; \quad \lim_{x \rightarrow \infty} \frac{\psi_R(x)}{x} = 0$$

Today's ingredients

Distribution of $\Lambda\mathbb{U}$

Supported by the ellipsoid $\{\Sigma u : u \in \mathcal{S}^{d-1}\}$ where

$$\Sigma := \Lambda' \Lambda$$

is (up to a constant) the covariance matrix of \mathbb{X}

The ellipsoid is centered at the origin and has d axes directed by the eigenvectors of Σ with semi-length the square roots of the corresponding eigenvalues

$$\lambda_1 = \dots = \lambda_{\mathbf{m}} > \lambda_{\mathbf{m}+1} \geq \dots \geq \lambda_d > 0$$

ordered and repeated, where $1 \leq \mathbf{m} \leq d$ is the multiplicity of the largest one. If $\mathbf{m} = d$ we have a spherical distribution.

Today's ingredients

Distribution of $\Lambda\mathbb{U}$

Up to an orthogonal transformation we may assume that

$$\Lambda = \text{diag}(\sqrt{\lambda_1}, \sqrt{\lambda_2}, \dots, \sqrt{\lambda_d})$$

Localization principle for R

If \mathbb{V} is a bounded random variable then the vector $R\mathbb{V}$ has a large norm iff R is large and \mathbb{V} is close to its maximum.

Therefore, when $\mathbb{X} = R\Lambda\mathbb{U}$ is large then $\|\mathbb{X}\|$ is of order $\sqrt{\lambda_1}R$ and \mathbb{X} is located near the dominant eigenspace associated with λ_1 :

$$\|\mathbb{X}\| = \sqrt{\lambda_1}R \left(1 - \sum_{k=m+1}^d \frac{\lambda_1 - \lambda_k}{\lambda_1} U_k^2 \right)^{1/2}$$

Starter

Theorem [FDS, 2015]

Define the functions ψ and ϕ on $(0, \infty)$ by

$$\psi(x) = \sqrt{\lambda_1} \psi_R\left(\frac{x}{\sqrt{\lambda_1}}\right) \quad \text{and} \quad \phi(x) = \left(\frac{\psi(x)}{x}\right)^{1/2}$$

Then, as $x \rightarrow \infty$,

$$\mathbb{P}(\|\mathbb{X}\| > x) \sim C_{\mathbf{m}} (\phi(x))^{d-\mathbf{m}} \mathbb{P}(R > \frac{x}{\sqrt{\lambda_1}})$$

where

$$C_{\mathbf{m}} := \frac{\Gamma(\frac{d}{2})}{\Gamma(\frac{\mathbf{m}}{2})} 2^{(d-\mathbf{m})/2} \left(\prod_{k=\mathbf{m}+1}^d \frac{\lambda_1}{\lambda_1 - \lambda_k} \right)^{1/2}$$

In particular, $\|\mathbb{X}\|$ is also in the max domain of attraction of the Gumbel distribution

Starter

Theorem [FDS, 2015]

Define $\Theta = \frac{1}{\|\mathbb{X}\|} \mathbb{X} = (\Theta_1, \dots, \Theta_d)$.

Then, as $x \rightarrow \infty$, conditionally on $\|\mathbb{X}\| > x$,

$$\left(\frac{\|\mathbb{X}\| - x}{\psi(x)}, \Theta_1, \dots, \Theta_m, \frac{\Theta_{m+1}}{\phi(x)}, \dots, \frac{\Theta_d}{\phi(x)} \right)$$

converges in distribution to

$$\left(E, \Theta^{(m)}, \sqrt{\frac{\lambda_{m+1}}{\lambda_1 - \lambda_{m+1}}} G_{m+1}, \dots, \sqrt{\frac{\lambda_d}{\lambda_1 - \lambda_d}} G_d \right)$$

where E is an exponential random variable with mean 1, $\Theta^{(m)}$ is uniformly distributed on \mathcal{S}^{m-1} , G_{m+1}, \dots, G_d are independent standard Gaussian random variables, and all components are independent.

Starter

$\|\mathbb{X}\|$ is in the max domain of attraction of the Gumbel distribution

Thus :

- Consider $a_n > 0$ such that $\mathbb{P}(\|\mathbb{X}\| > a_n) \sim \frac{1}{n}$
- Set $b_n = \psi(a_n)$

Then :

- $\lim_{n \rightarrow \infty} a_n = \infty$ and $\lim_{n \rightarrow \infty} \frac{b_n}{a_n} = 0$
- For all $t \in \mathbb{R}$,

$$\lim_{n \rightarrow \infty} \mathbb{P} \left(\frac{M_n - a_n}{b_n} \leq t \right) = e^{-e^{-t}}$$

Starter

Corollary [FDS, 2015]

Let $c_n = \phi(a_n)$ and define the points

$$P_{n,i} = \left(\frac{\|\mathbb{X}_i\| - a_n}{b_n}, \Theta_{i,1}, \dots, \Theta_{i,m}, \frac{\Theta_{i,m+1}}{c_n}, \dots, \frac{\Theta_{i,d}}{c_n} \right)$$

Then, the point processes $\sum_{i=1}^n \delta_{P_{n,i}}$ converge weakly to a PPP $\sum_{i=1}^{\infty} \delta_{P_i}$ on $\mathbb{R} \times \mathcal{S}^{m-1} \times \mathbb{R}^{d-m}$ with

$$P_i = \left(\Gamma_i, \Theta_i^{(m)}, \sqrt{\frac{\lambda_{m+1}}{\lambda_1 - \lambda_{m+1}}} G_{i,m+1}, \dots, \sqrt{\frac{\lambda_d}{\lambda_1 - \lambda_d}} G_{i,d} \right)$$

where $\{\Gamma_i\}$ are the points of a PPP on $(-\infty, \infty]$ with mean measure $e^{-t} dt$, $\{\Theta_i^{(m)}\}$ are i.i.d. vectors uniformly distributed on \mathcal{S}^{m-1} and $\{G_{i,k}\}$ are i.i.d. standard Gaussian variables, all sequences being mutually independent.

Conclusion

Vectors $\mathbb{X}_i = \|\mathbb{X}_i\|\Theta_i$ with the largest norm concentrate around the dominant eigenspace in such a way that

- $\|\mathbb{X}_i\| \sim a_n + b_n\Gamma_i$ with $a_n \rightarrow \infty$ and $b_n = o(a_n)$
- The \mathbf{m} first coordinates of Θ_i are uniform on $\mathcal{S}^{\mathbf{m}-1}$
- The $d - \mathbf{m}$ other coordinates of Θ_i tend to 0 with rate $c_n \rightarrow 0$ with Gaussian fluctuations

Last question

Are these large vectors always diametrically opposed?

- If $m = 1$: S^{m-1} has only one direction
Thus two vectors with a large norm will be on opposite sides and their distance is automatically large, typically twice as large as the norm of each one.
We expect that D_n behaves roughly like $2a_n$
- If $m > 1$: S^{m-1} has an infinite number of directions
Thus two vectors with a large norm can be close to each other and their distance will be typically much smaller than twice their norm.
We expect then a corrective term when comparing D_n to $2a_n$

Theorem [FDS, 2015]

Assume that $\mathbf{m} = 1$, i.e. $\lambda_1 > \lambda_2$.

Then

$$\frac{D_n - 2a_n}{b_n} \xrightarrow{(d)} \max_{i,j \geq 1} \left\{ \Gamma_i^+ + \Gamma_j^- - \frac{1}{4} \sum_{k=2}^d \frac{\lambda_k}{\lambda_1 - \lambda_k} (G_{i,k}^+ + G_{j,k}^-)^2 \right\}$$

where $\{\Gamma_i^\pm\}$ are the points of a PPP with mean measure $\frac{1}{2}e^{-t}dt$ on \mathbb{R} , and $\{G_{i,k}^\pm\}$ are i.i.d. standard Gaussian variables, independent of the points $\{\Gamma_i^\pm\}$.

Main course

Theorem [FDS, 2015]

Assume that $\mathbf{m} > 1$.

Then, for all $t \in \mathbb{R}$,

$$\lim_{n \rightarrow \infty} \mathbb{P} \left(\frac{D_n - 2a_n}{b_n} + d_n \leq t \right) = e^{-e^{-t}}$$

where

$$d_n = \frac{\mathbf{m} - 1}{2} \log \frac{a_n}{b_n} - \log \log \frac{a_n}{b_n} - \log C'_m$$

with

$$C'_m = (2d - \mathbf{m} - 1) 2^{\mathbf{m}-4} \pi^{-1/2} \Gamma\left(\frac{\mathbf{m}}{2}\right) \left(\prod_{k=\mathbf{m}+1}^d \frac{\lambda_1}{\lambda_1 - \lambda_k} \right)^{-1/2}$$

Example : bivariate Gaussian variable with correlation $\rho \in (0, 1)$

$$\mathbb{X} = R\Lambda\mathbf{U} \quad \text{with} \quad R = \sqrt{\chi_2^2} \quad \text{and} \quad \Lambda'\Lambda = \Sigma = \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}$$

- Eigenvalues : $\lambda_1 = 1 + \rho$ and $\lambda_2 = 1 - \rho$
- Eigenspaces : $\text{span}\{(1, 1)\}$ and $\text{span}\{(-1, 1)\}$
- Multiplicity : $\mathbf{m} = \begin{cases} 1 & \text{if } \rho \neq 0 \\ 2 & \text{if } \rho = 0 \end{cases}$

Main course

Set $a_n = \sqrt{(1 + \rho) \log n}$ and $b_n = \sqrt{\frac{1 + \rho}{2 \log n}}$

- If $\rho \neq 0$ then

$$\frac{D_n - 2a_n}{b_n} \xrightarrow{(d)} \max_{i,j \geq 1} \left\{ \Gamma_i^+ + \Gamma_j^- - \frac{1 - \rho}{8\rho} (G_i^+ + G_j^-)^2 \right\}$$

where $\{\Gamma_i^\pm\}$ are the points of a PPP with mean measure $\frac{1}{2}e^{-t}dt$ on \mathbb{R} , and $\{G_i^\pm\}$ are i.i.d. standard Gaussian variables, independent of the points $\{\Gamma_i^\pm\}$.

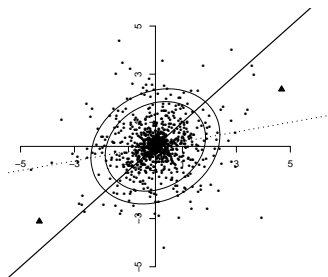
- If $\rho = 0$ then for all $t \in \mathbb{R}$,

$$\lim_{n \rightarrow \infty} \mathbb{P} \left(\frac{D_n - 2a_n}{b_n} + d_n \leq t \right) = e^{-e^{-t}}$$

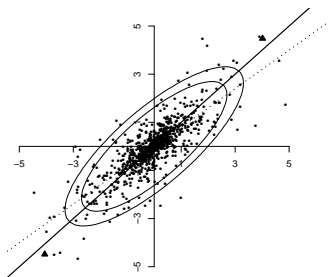
where

$$d_n = \frac{1}{2} \log \log n - \log \log \log n + \log(4\sqrt{2\pi})$$

Main course



$$\rho = 0.2$$



$$\rho = 0.8$$

The two points \blacktriangle realizing the diameter

- They concentrate around the diagonal at rate $O(\log n)$
- Fluctuations are Gaussian variables with variance $\frac{1-\rho}{2\rho}$

Possible generalizations thanks the localization principle

Distribution of \mathbb{X}

$$\mathbb{X} = R\lambda(\mathbb{U}) \text{ with } \lambda \text{ a bounded function}$$

The behavior of \mathbb{X} given that its norm is large and then the behavior of D_n will be determined by the maxima of $\|\lambda\|$:

- If they are isolated points, we obtain results similar to the case $\mathbf{m} = 1$
- Otherwise, if $\|\lambda\|$ is constant on non empty open subsets of \mathcal{S}^{d-1} , we obtain results similar to the case $\mathbf{m} > 1$

Possible generalizations thanks the localization principle

Non Euclidean diameter

Another open question in Jammalamadaka and Janson :

Asymptotic of the ℓ^q -diameter of a random spherical cloud ?

Consider :

- Spherical distribution : $\Lambda = I_d$ i.e. $\mathbb{X} := RU$
- The ℓ^q -diameter of the cloud :

$$D_n^{(q)} := \max_{1 \leq i < j \leq n} \|\mathbb{X}_i - \mathbb{X}_j\|_q$$

where, for $q \geq 1$, $\|x\|_q$ is the ℓ^q -norm of a vector $x \in \mathbb{R}^d$

Non Euclidean diameter

For $d \geq 2$ and $q \geq 1$, $q \neq 2$, the maximum of the ℓ^q -norm is achieved on the Euclidean sphere \mathcal{S}^{d-1} at isolated points :

- If $q \in [1, 2)$ then $\max_{u \in \mathcal{S}^{d-1}} \|u\|_q = d^{1/q-1/2}$ achieved at the 2^d diagonal points $(\pm d^{-1/2}, \dots, \pm d^{-1/2})$
- If $q \in (2, \infty)$, then $\max_{u \in \mathcal{S}^{d-1}} \|u\|_q = 1$ achieved at the $2d$ intersections of the axes with \mathcal{S}^{d-1}

Therefore the localization phenomenon will occur : a spherical vector \mathbb{X} such that $\|\mathbb{X}\|_q$ is large must be close to the direction of one of these maximum, and $D_n^{(q)}$ will be achieved by points which are nearly diametrically opposed along one of these directions.

Theorem [FDS, 2015]

If $q \in [1, 2)$, then

$$\frac{D_n^{(q)} - 2a_n^{(q)}}{b_n^{(q)}} \xrightarrow{(d)} \max_{1 \leq j \leq 2^{d-1}} \max_{i, i' \geq 1} \left\{ \Gamma_{i,j}^+ + \Gamma_{i',j}^- - \frac{q-1}{4} \sum_{k=1}^d (G_{i,j,k}^+ + G_{i',j,k}^-)^2 \right\}$$

where $\Gamma_{i,j}^\pm$ are the points of independent PPP on $(-\infty, \infty]$ with mean measure $2^{-d}e^{-t}dt$ and $(G_{i,j,1}^\pm, \dots, G_{i,j,d}^\pm)$ are i.i.d. Gaussian vectors with covariance matrix

$$\frac{1}{d(2-q)} \begin{pmatrix} d-1 & -1 & \dots & -1 \\ -1 & d-1 & \dots & -1 \\ \vdots & \vdots & \ddots & \vdots \\ -1 & \dots & -1 & d-1 \end{pmatrix}$$

Theorem [FDS, 2015]

If $q \in (2, \infty)$, then

$$\frac{D_n^{(q)} - 2a_n^{(q)}}{b_n^{(q)}} \xrightarrow{(d)} \max_{1 \leq i \leq d} \{\Gamma_i^+ + \Gamma_i^-\}$$

where $\{\Gamma_i^\pm\}$ are independent Gumbel random variables with location parameter $\log 2d$.

Thank you for your attention

Complete recipes in :

The diameter of an elliptical cloud

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Electron. Journal. Probab. 20 n°27, 1-32, 2015